

Research and Applications

A novel hyperparameter search approach for accuracy and simplicity in disease prediction risk scoring

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Abstract

Objective: Develop a novel technique to identify an optimal number of regression units corresponding to a single risk point, while creating risk scoring systems from logistic regression-based disease predictive models. The optimal value of this hyperparameter balances simplicity and accuracy, yielding risk scores of small scale and high accuracy for patient risk stratification.

Materials and Methods: The proposed technique applies an adapted line search across all potential hyperparameter values. Additionally, DeLong test is integrated to ensure the selected value produces an accuracy insignificantly different from the best achievable risk score accuracy. We assessed the approach through two case studies predicting diabetic retinopathy (DR) within six months and hip fracture readmissions (HFR) within 30 days, involving cohorts of 90 400 diabetic patients and 18 065 hip fracture patients.

Results: Our scores achieve accuracies insignificantly different from those obtained by existing approaches, reaching AUROCs of 0.803 and 0.645 for DR and HFR predictions, respectively. Regarding the scale, our scores ranged 0-53 for DR and 0-15 for HFR, while scores produced by existing methods frequently spanned hundreds or thousands.

Discussion: According to the assessment, our risk scores offer simple and accurate predictions for diseases. Furthermore, our new DR score provides a competitive alternative to state-of-the-art risk scores for DR, while our HFR case study presents the first risk score for this condition.

Conclusion: Our technique offers a generalizable framework for crafting precise risk scores of compact scales, addressing the demand for user-friendly and effective risk stratification tool in healthcare.

Key words: disease prediction; risk scoring system; hyperparameter search; electronic health record.

Introduction

Risk scoring systems have emerged as a favored approach to predict a range of health conditions in diverse healthcare settings. Notable examples include the Framingham Risk Scores^{1,2} and SCORE³ for foreseeing coronary heart disease, LACE⁴ and HOSPITAL⁵ for anticipating death or readmission after hospital discharge, IScore⁶ for predicting death and disability after an acute stroke, and Mortality Risk Score⁷ for estimating mortality in adults. These risk scoring systems often trace their development methodology back to the regression coefficient-based scoring principles.⁸ Building upon these foundational principles, Sullivan et al⁹ presented a comprehensive and systematic approach that has found significant traction in real-world healthcare scenarios and has been employed in creating well-known scoring systems such as the Framingham Risk Score and LACE.

The benefits of risk score systems are manifold. Firstly, they can provide clinicians with an easy-to-understand tool for estimating patient risk and making informed medical decisions.^{2,10} By utilizing score systems, healthcare professionals can assess the likelihood of specific health outcomes or complications, aiding in treatment plans and preventive measures.^{11,12} Additionally, a user-friendly risk score system also promotes patient engagement and behavior change. When patients understand their risk scores, they are more likely to comprehend potential health consequences, leading to active participation in health management and adopting beneficial lifestyle changes.⁹

Although the risk score system offers many advantages, little improvement has been made to the score derivation methodology since the earlier work performed by Sullivan et al.⁹ A notable gap pertains to a hyperparameter defined as *the*

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number of regression units in the disease prediction model to be mapped to a single point in the risk scoring system. A typical example of the “regression units” is the log-odds in the logistic regression model. For simplicity, we henceforth denoted this hyperparameter as B . Specifically, the gap is that effective approaches in determining a suitable value for B have not been adequately explored. The B value is important as it determines the granularity of the risk score. Higher granularity, corresponding to low B values, means using more risk score points to correspond to a given amount of regression-modeled risk. It results in a larger and more complex scale for the score system, which may introduce practical inconveniences in real-world implementation. On the plus side however, an intuitive benefit of the highly granular scoring system is that it captures a greater amount of information from the original regression model, consequently preserving better predictive accuracy, as measured by area under ROC curve (AUROC).¹³ When a scoring system has low granularity, intuitively, it sacrifices information from the regression model, thereby compromising accuracy. Nonetheless, relatively low granularity results in a smaller scale that finds widespread adoption in real-world healthcare settings due to its simplicity. Notable examples of risk scores with narrowed ranges that have gained significant usage in practical healthcare contexts include LACE⁴ and HOSPITAL.⁵ Therefore, addressing the challenge posed by scoring systems with either a low granularity, resulting in reduced predictive precision, or a highly granular scale leading to practical inconveniences, necessitates the development of a scoring approach that strikes a balance between the scale simplicity and the prediction accuracy. Although grid search, a classical hyperparameter tuning technique in machine learning, may be used to handle the issue, it can be computationally intensive when dealing with a multitude of hyperparameters, each having a wide range of possible values.¹⁴

In order to fill the gap, we have established two main objectives for this study: (1) Develop a novel hyperparameter search algorithm to identify the “best” amount of regression units in a disease prediction model, which should correspond to a single point in a risk scoring system for achieving a balance between the scale and accuracy for the risk score. (2) Assess the algorithm’s ability to generate compact-scale risk scores that preserve the majority of predictive accuracy from the root regression models by conducting two case studies, one on predicting diabetic retinopathy (DR) and the other on predicting hip fracture readmission (HFR).

Methods

Data source and preprocessing

In this study, we utilized the Oracle Cerner Health Facts Electronic Health Records (EHR) data warehouse as our data source. Health Facts comprises clinical data extracted from over 200 hospitals across the United States that operate on Cerner EHR systems during 2000-2018. The data encompasses a wide range of information, including patients’ time-stamped encounters, demographics, diagnoses, procedures, medications, laboratory results, vital signs, etc. Oracle Cerner collects and integrates the data in accordance with established procedures that adhere to the Health Insurance Portability and Accountability Act (HIPAA) laws. The Institutional Review Boards (IRB) at Oklahoma State

University (OSU) exempted the study from review because the data has been completely de-identified according to HIPAA regulations. All the data collection, preprocessing, and analysis involved in this study were performed on the devices hosted at OSU.

Our two case studies involved leveraging large-scale EHR datasets from Health Facts to predict DR and HFR. DR is a complication of diabetes that can cause vision loss or blindness over time if not diagnosed early enough and left untreated.^{15,16} Hip fractures (HF) significantly increase morbidity and mortality in older adults, frequently resulting in post-discharge readmissions.^{17,18} Both are significant conditions drawing extensive research attention and warranting further investigation.

- **DR Data:** We extracted the DR case and control cohorts using the same diagnosis codes and a similar cohort derivation method as utilized in a prior study by Wang et al.¹⁹ Together with the cohorts, we gathered 31 variables related to patient’s demographics, duration of diabetes, complications, and laboratory results, all of which have been shown in the literature to be significantly associated with DR.²⁰⁻²⁴ In the predictive modeling, the values of these variables, during a two-year window that was 6 months preceding the first diagnosis of DR, were averaged to predict whether DR would occur within the 6-month period. This approach models the DR prediction in six months given a diabetic encounter and history in past two years.^{19,25} Subsequently, we applied the complete-case preprocessing method to these variables. After preprocessing, the variables maintained distributions close to those of the raw data (the distribution plots are available in Section B of the [Supplementary Material](#)). Next, we applied a machine learning-based ensemble predictor selection method²⁶ to identify a reduced set of key predictors from the 31 variables. These key predictors enabled us to create a concise yet accurate risk scoring system.
- **HFR Data:** Regarding the selection of the patient cohort for HFR, we extracted data from Health Facts and followed a cohort derivation method similar to that used in a prior HFR study.²⁷ The data comprised patient demographics, historical visits, diagnoses, procedures, and seven laboratory results, constituting the initial set of variables for our analysis. To maintain methodological consistency across both case studies, we applied the same complete-case preprocessing and predictor selection methods used for DR data to this HFR dataset as well. Beyond the processing, the selected key predictors were utilized to predict all-cause readmissions within 30 days from the HF inpatient visits.

Detailed diagnosis and procedure codes for patient extraction, flow charts outlining data preprocessing, and initial sets of variables for analysis for the two case studies are provided in Section A of the [Supplementary Material](#). For both study cohorts, we randomly partition the data into training (70%) and test subsets (30%) for predictive analysis.

Risk score derivation methods

Figure 1 shows a risk scoring framework adapted from the one established by Sullivan et al.,⁹ serving as the foundational pipeline for our risk score derivation. Our novel

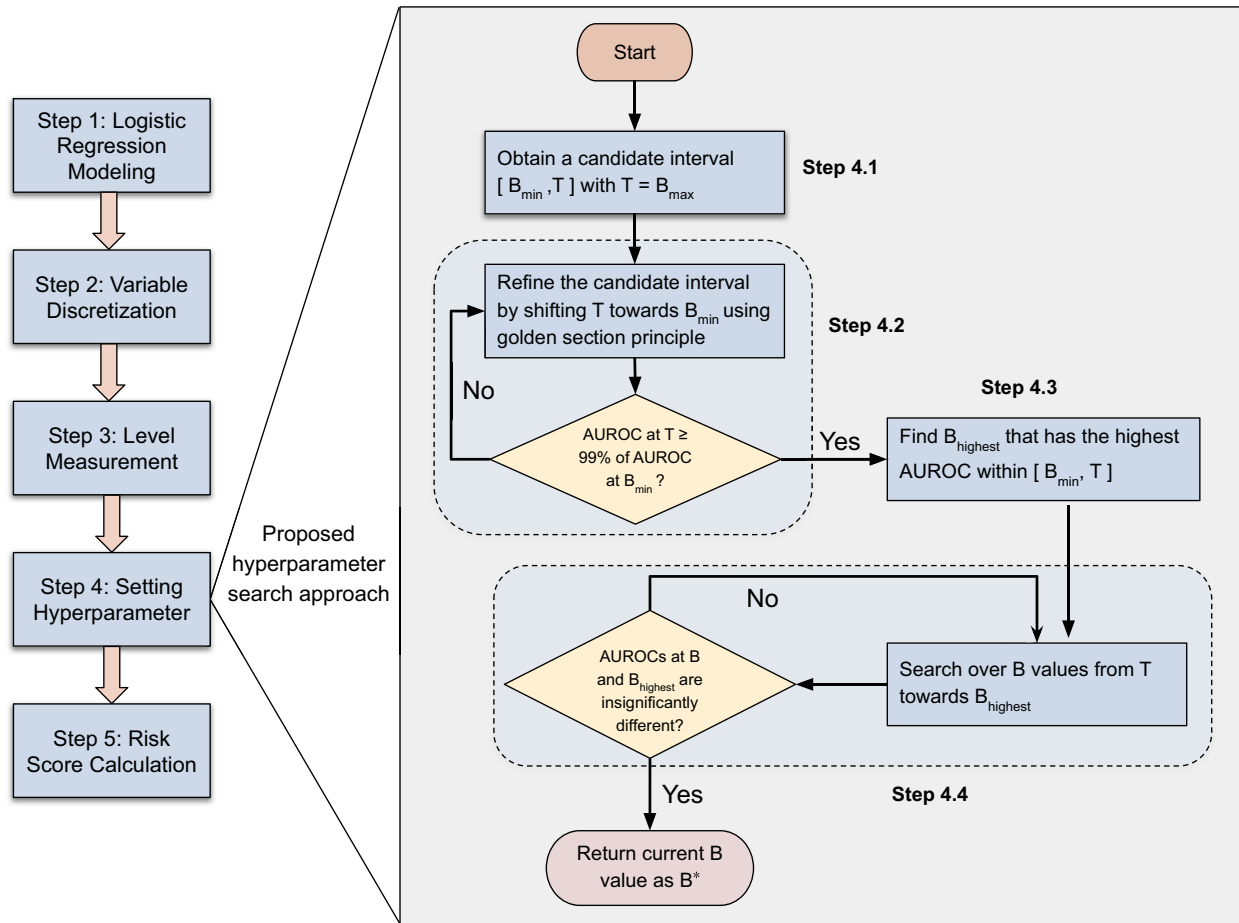


Figure 1. Flowchart illustrating the risk score derivation framework and our refinements in Step 4. In this illustration, B denotes the number of regression units in the disease prediction model to be mapped to a single point in the risk scoring system. T is the right end of the candidate interval of B values, and updated iteratively in the algorithm.

hyperparameter search approach centers on the step of *Setting Hyperparameter B*, aiming to develop simple yet accurate risk scores, as detailed in the following.

Scoring framework

Step 1. Logistic Regression Modeling: Construct a logistic regression model to predict the presence of a health condition (modeled as a binary target variable y) based on n predictors, denoted by x_1, x_2, \dots, x_n . The model can be represented as Equation (1):

$$\ln \frac{p}{1-p} = \beta_0 + \sum_{i=1}^n \beta_i x_i, \tag{1}$$

where p represents the probability that $y=1$, indicating that patients developed DR or were readmitted respectively in our two case studies. The value $\ln \frac{p}{1-p}$, known as “log-odds,” is used to model patient risk of having the health condition. While β_0 is the intercept and β_i represents the coefficient for the predictor x_i .

Step 2. Variable Discretization: Convert continuous predictors to categorical variables by discretizing them into multiple intervals (aka the levels of the resulted

ordinal categorical variables) using meaningful cut-offs based on medical expertise. Statistical methods, such as percentile-based cutoffs, are frequently employed to discretize continuous variables in the literature.¹² Our comparison (provided in Section C of the [Supplementary Material](#)) revealed minor differences in AUROC between the medically meaningful and statistical cutoffs. Hence, in this article, we focus on reporting the results based on medically meaningful discretization due to the clinical relevance and interoperability of this approach.

Step 3. Regression Unit Measurement for Levels: This step involves measuring the regression units, specifically log-odds in our case, for every level of each categorical variable. The measurement follows the subsequent procedure: Given any variable x_i , we first determine a reference value for each level, which is the mid-value for intervals and a modeled value in logistic regression for categorical variables (eg 0 for modeling female and 1 for modeling male). Then, the level with the lowest reference value corresponds to the lowest-risk level if the coefficient is positive; otherwise, it is the level with the highest reference value. Denote the reference value of the lowest-risk level as

W_{min} , then for a level of x_i with reference value of W , the log-odds assigned to the level are expressed as $\beta_i(W - W_{min})$.

- Step 4.** Setting Hyperparameter B : The hyperparameter, in this case, is the number of log-odds corresponding to a single risk score point. It can be determined by multiplying the coefficient of a selected base variable by a factor, such as $\beta_{age} \times 5$. However, the approaches for selecting the base variable and the factor in current literature lack a delicate design to generate risk scores that achieve both high simplicity and accuracy. Our novel hyperparameter search algorithm, elaborated in the next subsection on “Hyperparameter Searching,” addresses this gap, constituting the primary innovation of this study.
- Step 5.** Risk Score Calculation: Once B is determined, the associated risk score for each level of a predictor can be calculated using the formula $\beta_i(W - W_{min})/B$ and round it to the nearest integer. The overall risk score of a patient will be the sum of the risk scores corresponding to each variable’s measurement of the patient.

Hyperparameter searching

As discussed in the Introduction section, a smaller value for the hyperparameter B leads to a more granular risk score, preserving greater predictive power from the regression model. However, it may result in an unnecessarily large scale, posing inconvenience for clinical applications. On the other hand, a larger B value yields a simpler scale but incurs a loss of accuracy. Hence, a clever choice of the B value is crucial for simplifying the risk score system without compromising accuracy. Many scores used a multiple of β_{age} ,^{4,6,7,9} to account for increasing risk associated with aging, while some other studies employed the smallest coefficient^{5,12} to ensure all scores to be larger than one. However, none of the approaches adequately considered both the scale and accuracy of the risk score. Grid-search-based enumeration across all predictors and all potential factor values for each predictor is an intuitive approach to tackle the issue, but it can be computationally expensive and time-consuming.¹⁴ Our new approach, rather than engaging in a two-dimensional search across variables and factors, executes a uni-dimensional search directly over all feasible B values. The flow diagram is illustrated in Figure 1, with steps explained below:

- Step 4.1** Obtain all possible B values by multiplying the coefficient of each variable by all potential factor values (we used 1, 2, ..., 10 in our implementation). Then, sort the resulting B values in an ascending order and define a candidate interval $[B_{min}, T]$ with $T = B_{max}$ initially to cover the entire range of B values.
- Step 4.2** Iteratively refine the candidate interval by adjusting the right endpoint T from B_{max} towards B_{min} until the accuracy at T reaches at least 99% of the accuracy at B_{min} . In our implementation, we measure accuracy using AUROC. The endpoint adjustment adheres to the golden section principle.²⁸ In other words, for each iteration, the new value of T is updated as $T' - 0.382 \times (T' - B_{min})$, where T' represents the previous value of T .

Step 4.3 Within the refined candidate interval $[B_{min}, T]$, identify the B value associated with the highest AUROC, denoted as $B_{highest}$.

Step 4.4 Search from the right endpoint of the refined candidate interval T towards $B_{highest}$ to find the first B value whose AUROC is insignificantly different from that of $B_{highest}$ via DeLong test at a 0.01 significance level.²⁹ Finally, return the found B value, denoted as B^* .

The benefit of this search strategy is that we can leverage the intuition that with the increase of B , the AUROC demonstrates an overall declining trend as larger B tends to yield less granularity in the risk score. The trend enables us to perform directional search to find a suitable B value sooner. Specifically, Step 4.2 enables us to quickly skip B values close to the right end of the trend that are associated with low accuracies, as illustrated in Figure 2. Furthermore, once the refined candidate interval is determined, the search from T to $B_{highest}$, as described in Step 4.4, saves effort of performing DeLong test exhaustively for the B values less than $B_{highest}$.

All the data cleaning, analysis and algorithm development presented in this article were implemented using Python 3.10. The logistic regression models used in this study were created and executed using the “*glm()*” function from the Python *statsmodels* 0.14.0 module. Our code is publicly available on GitHub at <https://github.com/yajun668/RiskScoring>.

Results

Descriptive statistics of case study cohorts

After preprocessing, our DR cohort consisted of 90 400 diabetic patients, among whom 3380 were diagnosed with DR. The HFR cohort included 18 065 HF patients, among whom 2055 were readmitted to the hospital within 30 days from their HF inpatient visits. The selected key predictors and their detailed statistics within the training and test datasets of the two cohorts are presented in Table 1, showing insignificant difference across all variables except for creatinine between the training and test datasets.

Trend between AUROC and B

Figure 3A and C depicts the relationships between AUROC and B values for DR and HFR predictions, respectively. Both plots demonstrate a consistent downward trend, aligning with the intuitive expectation that higher B values lead to lower granularity of the risk scoring system, ultimately compromising its accuracy. Figure 3B and D provides zoomed-in views of the refined candidate intervals, showing that $B_{highest}$ does not necessarily coincide with the smallest B . Furthermore, many AUROCs in the interval appear very close, indicating that towards the right-hand side of the interval, there are competitive B values that could result in narrower scales of risk scores, with statistically insignificant differences in accuracy compared to that at $B_{highest}$. All the observations favorably support the design of our proposed hyperparameter search algorithm.

Score system comparison

To assess the effectiveness of our proposed approach, we compared the risk scores developed using B^* with those derived based on $5\beta_{age}$ and B_{min} —two commonly used values

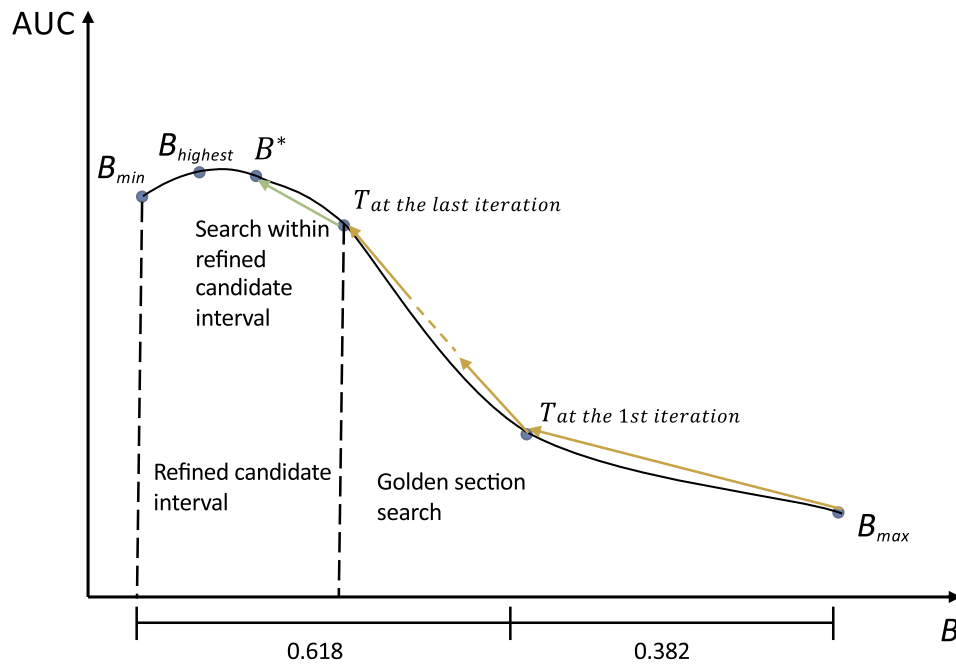


Figure 2. An illustration of the search trajectory of the developed hyperparameter search algorithm.

for the hyperparameter B in the literature. Table 2 summarizes the AUROCs and scales of the risk scores across different B values. Upon observation, our algorithm consistently generates risk scores closely aligned with those obtained using other B values in terms of AUROC. It is worth noting that, though the AUROC associated with B^* is slightly lower than those associated with other B values for HFR, the difference is statistically insignificant according to DeLong test. Remarkably, the scales of the risk scores derived through our approach exhibit much simpler ranges. In the context of DR prediction, our risk score ranged up to only 53, in contrast to 191 for $5\beta_{age}$ and 11 018 for B_{min} . Similarly, for HFR prediction, our highest score is only 15, whereas $5\beta_{age}$ and B_{min} have ranges with the highest score as high as 93 and 18 728, respectively.

We additionally compared the risk scores with the corresponding logistic regression models—the root model from which the scores are derived—in terms of AUROC. The AUROC plots are displayed in Figure 4, illustrating marginal differences, up to 0.029 (as observed for $5\beta_{age}$ for DR), between the predictive accuracy achieved by the risk scores and that of logistic regressions. This aligns with what has been reported in the literature,¹⁹ reiterating the effectiveness of the entire risk scoring framework in maintaining strong predictive capacity from logistic regressions. Figure 4 also shows comparable area under the precision-recall curve (AUPRC)³⁰ across evaluated risk scores for each condition. They all outperformed random models, suggesting their ability in differentiating patients by disease risk. DR risk scores outperformed the random model to a greater extent than HFR risk scores, which aligns with AUROC findings. However, the class imbalance in our data likely limited AUPRC performance. Integrating techniques for handling imbalanced

data into regression models has the potential to improve AUPRC for risk scores.

Furthermore, we report the new risk score systems for DR and HFR in Table 3. Note that the risk score derived using $5\beta_{age}$ for DR, presented in the table, is essentially a variation of the score system proposed by Wang et al¹⁹ with two additional predictors and slightly adjusted scores for certain levels. Compared to it, the risk score derived using our approach, B^* , significantly simplified the system, by aggregating many levels across a multitude of predictors, such as <60 and [60,80) for glucose, as well as “African American” and “Other” for race. These levels can be combined because they share the identical risk points. A similar finding can be observed for the HFR risk score as well. Many levels can be combined, for instance, <65 and [65,75) for age, and <9 and [9,9.7) for hemoglobin. More interestingly, our new scoring system requires an even more concise set of nine predictors, specifically BUN, hemoglobin, hematocrit, length of stay, preInp1Y, preER1Y, Charlson comorbidity index, age, and platelet count, rather than the 12 variables chosen by feature selection, because the other predictors exhibit 0 risk points across all levels, resulting in no effect on final risk score.

Discussion

The widespread deployment of EHR systems has made a tremendous volume of digitized clinical data available. Coupled with advancements in medical informatics and analytics, it has provided valuable and actionable insights for addressing a wide range of healthcare challenges, including the high-cost patients identification, disease prediction, patient triaging, and treatment plan optimization, among others.^{31–34} Machine learning and deep learning models are often employed to tackle the challenges because of their high

Table 1. Descriptive statistics on training and test datasets for DR and HFR predictions.

	DR dataset					
	Training		Test		P-value ^a	
	Non-DR	DR	Non-DR	DR	Non-DR	DR
# Patient (%)	60 936 (96.3)	2344 (3.7)	26 084 (96.2)	1036 (3.8)	–	–
Creatinine, mean (SD)	1.06 (0.45)	1.96 (1.88)	1.06 (0.45)	1.80 (1.60)	0.355	0.018
HbA1c, mean (SD)	7.13 (1.50)	8.35 (2.03)	7.14 (1.51)	8.39 (1.97)	0.173	0.605
Diabetes duration, mean (SD)	1.92 (1.76)	2.75 (2.02)	1.92 (1.77)	2.76 (2.05)	0.999	0.929
White blood cell, mean (SD)	8.11 (2.19)	7.97 (2.21)	8.12 (2.19)	8.01 (2.31)	0.547	0.619
Glucose, mean (SD)	142.42 (46.17)	173.96 (61.61)	142.61 (46.24)	174.75 (62.36)	0.575	0.733
Age, mean (SD)	64.16 (14.08)	60.47 (13.37)	64.05 (14.07)	60.82 (12.77)	0.306	0.475
Hematocrit, mean (SD)	38.99 (4.71)	36.23 (4.72)	39.04 (4.69)	36.42 (4.71)	0.233	0.294
Sodium, mean (SD)	138.87 (2.46)	138.59 (2.37)	138.86 (2.45)	138.48 (2.39)	0.537	0.249
BUN, mean (SD)	19.66 (9.45)	27.45 (14.78)	19.67 (9.57)	26.82 (14.51)	0.846	0.245
Anion gap, mean (SD)	9.47 (2.55)	9.52 (2.71)	9.45 (2.55)	9.36 (2.62)	0.419	0.131
Nephropathy = yes (%)	3030 (5.0)	656 (28.0)	1281 (4.9)	278 (26.8)	0.715	0.516
Neuropathy = yes (%)	5197 (8.5)	782 (33.4)	2190 (8.4)	349 (33.7)	0.529	0.884
Race (%)						
African American	10 867 (17.8)	897 (38.3)	4669 (17.9)	377 (36.4)		
Caucasian	45 355 (74.4)	1270 (54.2)	19 332 (74.1)	586 (56.6)	0.417	0.435
Other	4714 (7.7)	177 (7.6)	2083 (8.0)	73 (7.0)		
	HFR dataset					
	Training		Test		P-value ^a	
	Non-HFR	HFR	Non-HFR	HFR	Non-HFR	HFR
# Patient	11 226 (88.78)	1419 (11.22)	4784 (88.27)	636 (11.73)	–	–
Age, mean (SD)	80.01 (9.80)	80.87 (9.22)	80.05 (9.74)	81.26 (9.04)	0.823	0.373
Length of stay, mean (SD)	5.32 (2.79)	6.18 (3.53)	5.34 (2.78)	6.27 (3.36)	0.819	0.607
Platelet count, mean (SD)	209.64 (81.53)	216.67 (89.54)	210.16 (82.10)	219.50 (91.35)	0.711	0.511
BUN, mean (SD)	19.53 (10.97)	24.01 (14.37)	19.55 (11.25)	23.87 (13.29)	0.904	0.835
Hemoglobin, mean (SD)	10.11 (1.34)	10.07 (1.31)	10.14 (1.35)	10.10 (1.29)	0.115	0.640
Creatinine, mean (SD)	0.97 (0.61)	1.15 (0.80)	0.97 (0.60)	1.11 (0.74)	0.748	0.307
Hematocrit, mean (SD)	30.06 (3.84)	30.09 (3.86)	30.17 (3.86)	30.12 (3.85)	0.093	0.860
CCI, mean (SD)	1.29 (1.48)	1.72 (1.61)	1.30 (1.48)	1.67 (1.61)	0.595	0.541
Potassium	4.02 (0.44)	4.06 (0.46)	4.01 (0.43)	4.04 (0.47)	0.200	0.452
Sodium	137.09 (3.74)	137.23 (4.03)	137.06 (3.77)	137.20 (3.81)	0.617	0.852
preInp1Y ^b (%)						
0	8335 (74.2)	926 (65.3)	3571 (74.6)	389 (61.2)		
1	1809 (16.1)	244 (17.2)	745 (15.6)	131 (20.6)	0.684	0.132
2	1082 (9.6)	249 (17.5)	468 (9.8)	116 (18.2)		
preER1Y ^c (%)						
0	7230 (64.4)	792 (55.8)	3068 (64.1)	321 (50.5)		
1	2127 (18.9)	288 (20.3)	906 (18.9)	152 (23.9)	0.905	0.063
2	1869 (16.6)	339 (23.9)	810 (16.9)	163 (25.6)		

Abbreviation: CCI = Charlson Comorbidity Index.

^a The P-values are associated with the statistical tests comparing variable differences between the training and test datasets.^b Number of inpatient visits within 1 year before.^c Number of emergency department visits within 1 year before.

predictive accuracy.^{35–37} However, the inherent black-box nature of machine/deep learning often poses challenges in interpreting the results for clinicians.³⁸ Additionally, many existing EHR systems in hospitals lack support for complex machine-learning models.³⁹

In contrast, risk scores are easy to interpret, understand, and implement in healthcare settings, contributing to their considerable attention and real-world applications. The novel hyperparameter search algorithm developed in this study

enable the creation of simple yet accurate risk scores, which can support medical decision making in various aspects of patient care. Firstly, risk scoring systems developed using comprehensive socioeconomic and clinical determinants enable healthcare professionals to compute patients' risk of developing specific conditions in the future. A high risk score can serve as early-warning tool, prompting timely intervention for effective care management. Additionally, the risk score's interpretability, along with insights into how each

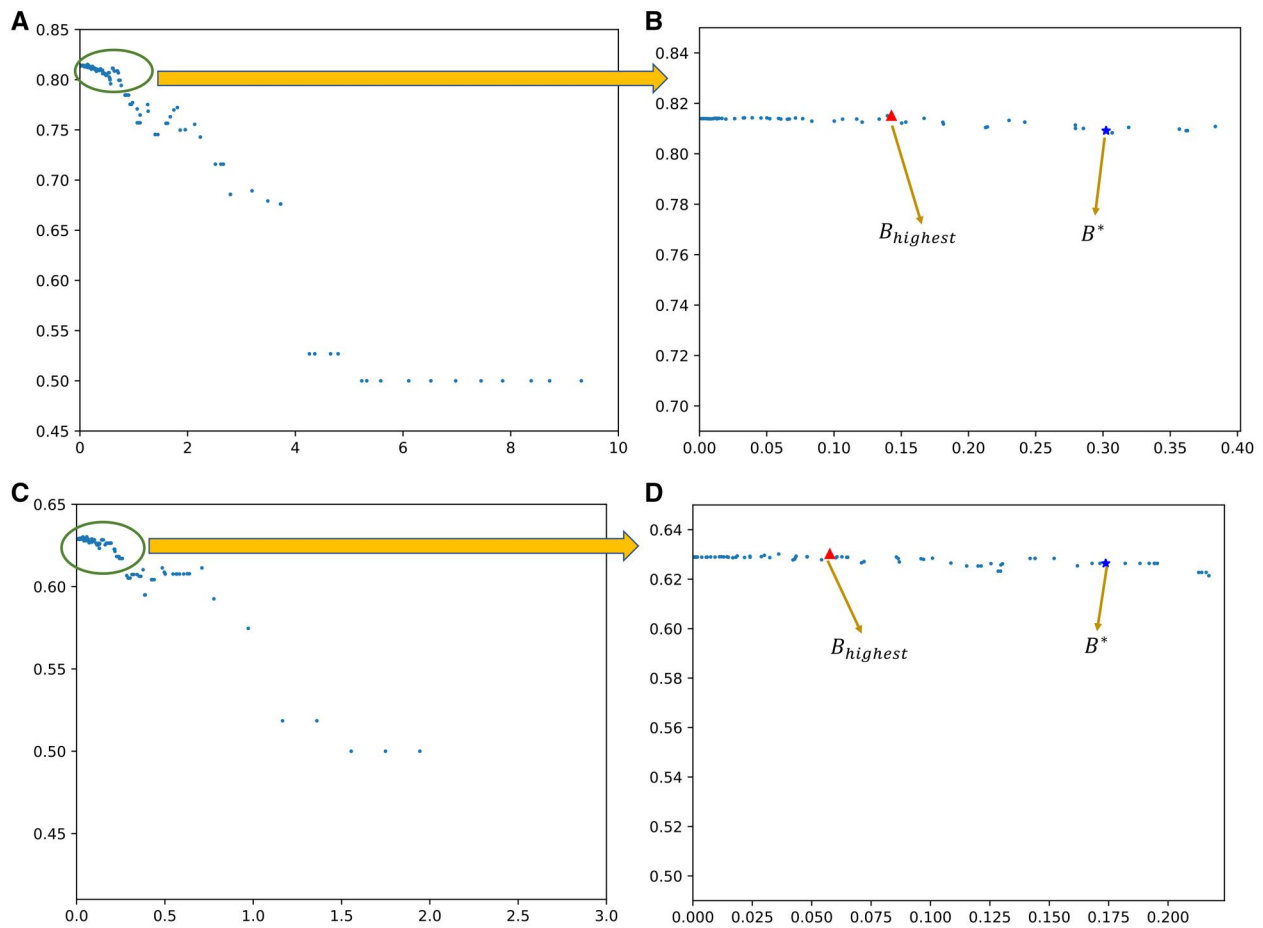


Figure 3. The AUROC and B relationships for DR (A and B) and HFR (C and D). (A) and (C) provide overall trends over all B values considered. (B) and (D) provide zoomed-in views of the AUROC- B relationship for DR and HFR, respectively.

Table 2. Comparison of AUROCs and risk score scales between the proposed method (B^*) and existing approaches (B_{\min} and $5\beta_{age}$) for DR and HFR.

	AUROC			Score scale		
	B^*	$5\beta_{age}$	B_{\min}	B^*	$5\beta_{age}$	B_{\min}
DR	0.803	0.802	0.804	0-53	0-191	0-11 018
HFR	0.645	0.651	0.652	0-15	0-93	0-18 728

feature contributes to the total risk score, empowers patients to better grasp the factors that pose health risks. With the knowledge, patients are more likely to take action to address the factors that negatively impact their health.^{40,41}

Compared to the state-of-the-art DR risk score,¹⁹ our new DR risk score system, generated using the algorithm proposed in this study, exhibits equivalently high accuracy with a significantly simpler scale. As for the risk score for HFR, to the best of our knowledge, this is the first study in developing a risk score system for this condition. The two new risk score systems not only demonstrate the effectiveness of our proposed approach but also offer highly potential alternatives, once externally validated, for the prediction and risk stratification for DR and HFR respectively. Furthermore, while our case studies concentrated solely on two conditions,

DR and HFR, our approach can serve as a general framework for developing risk scores for other health conditions as well.

Recommendations for accurate risk scoring

Our technique enables the creation of concise risk scores closely mirroring the accuracy of regression models. Therefore, robust regression models are the cornerstone for accurate risk scoring. Various factors spanning data collection, preprocessing, modeling, and deployment influence regression modeling, subsequently the accuracy of risk score, in real-world disease prediction applications. Key considerations encompass data representativeness, incorporation of comprehensive socioeconomic and clinical variables, handling missing values, addressing data imbalance, and the geographical and care setting differences between modeling and deployment. Analysts should select data from sources aligned with the geographic and care settings where the model will be deployed and thoroughly evaluate the data quality before modeling to ensure the relevance and robustness of their models in the target settings.^{42,43} Various methods exist for handling missingness and imbalance within health data, yet there is a lack of widespread consensus and acceptance within the scientific community regarding the most effective methodology.^{44,45} Distinct methodologies yield different models and predictive outcomes.⁴⁶ Analysts should carefully consider

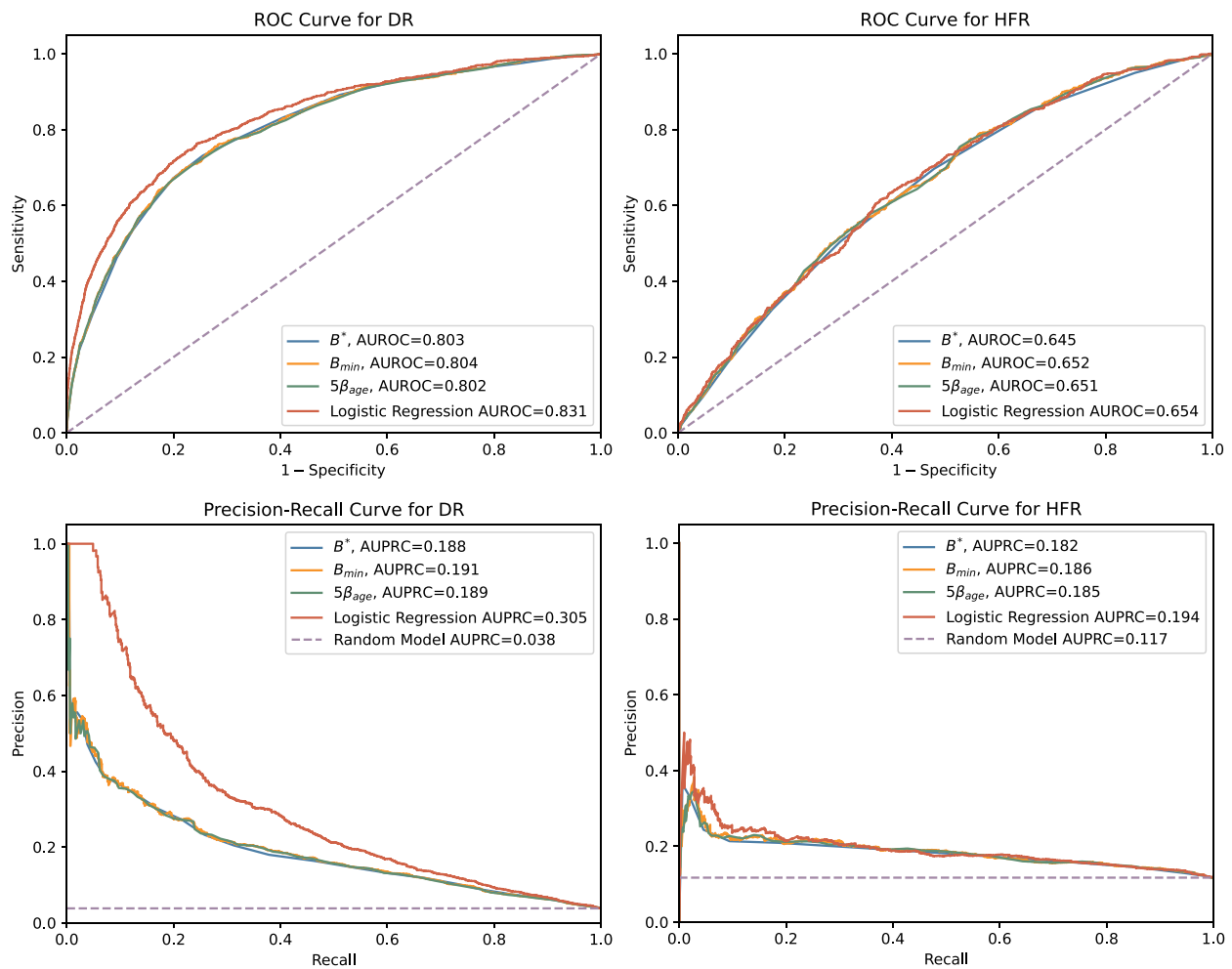


Figure 4. Comparison among risk scores and logistic regressions in AUROC and AUPRC.

various aspects, such as clinical relevance, interpretability, generalizability, and accuracy, for optimal model selection. When incorporating variables, analysts should thoroughly examine available metrics concerning both health and socioeconomics, then leverage clinical and biological expertise, along with feature selection techniques,⁴⁷ to identify essential predictors for accurate modeling. EHR data is rich in clinical data but often lacks socioeconomic variables, which are crucial for understanding and addressing many health conditions.^{48–52} By integrating socioeconomic factors into EHR,^{53,54} the endeavor of crafting more comprehensive and accurate risk scores can be significantly bolstered.

Limitations

There are several limitations with this study. (1) Our risk scoring algorithm is essentially a statistical approach revealing associations rather than establishing causality. The generated risk scores should be viewed as decision support tools for healthcare professionals, with application and interpretation contingent upon clinical expertise. (2) In the case studies, we optimized risk scores for high AUROCs. Other accuracy measures were not necessarily preserved to the same degree. (3) Socioeconomic factors were unavailable within the EHRs

used, thus omitted from the case studies. (4) The risk scores for DR and HFR created in the case studies require validation using external data.

Conclusion

In this study, we introduce a novel hyperparameter search algorithm intended to automatically determine an optimal amount of log-odds that should be calibrated to a single score in a risk scoring system to achieve a balance between accuracy and simplicity within the risk scoring system. The implications of our proposed approach in healthcare settings are substantial as it delivers simple yet accurate risk scores that support healthcare professionals and decision makers in patient stratification, treatment planning, and various medical decision-making processes. Additionally, on the patient side, the risk score encourages them to adopt healthier behaviors, undergo early screenings, and prioritize preventive measures before conditions deteriorate. Our future research will focus on evaluating the developed approach across a broader spectrum of health conditions and conducting external validations for our new DR and HFR risk score systems. In addition, exploring improved algorithms that can balance

Table 3. Risk scoring systems derived using B^* , $5\beta_{age}$, and B_{min} for DR and HFR.

Variable	Level	Risk score for DR		
		B^*	$5\beta_{age}$	B_{min}
Neuropathy	No	0	0	0
	Yes	3	11	638
Nephropathy	No	0	0	0
	Yes	2	6	365
Creatinine	<0.5	0	0	0
	[0.5, 1)	1	4	203
	[1, 1.5)	2	9	502
	[1.5, 2)	4	14	801
	≥ 2	7	24	1357
HbA1c	<6	0	0	0
	[6, 8)	2	7	383
	[8, 10)	4	13	767
	[10, 12)	6	20	1150
	≥ 12	8	30	1725
Diabetes duration	<1	0	0	0
	[1, 2)	0	2	96
	[2, 3)	1	3	192
	[3, 4)	1	5	287
	≥ 4	4	15	833
White blood cell	<4	4	14	773
	[4, 6)	3	12	694
	[6, 8)	3	10	589
	[8, 12)	2	8	431
	≥ 12	0	0	0
Glucose	<60	0	0	0
	[60, 80)	0	1	76
	[80, 100)	1	3	166
	[100, 200)	2	8	434
	≥ 200	7	24	1393
Age	<35	3	12	704
	[35, 50)	2	9	515
	[50, 65)	2	6	343
	[65, 75)	1	4	200
	[75, 85)	0	2	86
	≥ 85	0	0	0
Hematocrit	<30	6	21	1215
	[30, 35)	5	16	933
	[35, 40)	4	13	725
	[40, 50)	2	7	415
	≥ 50	0	0	0
Sodium	<136	0	0	0
	[136, 144)	3	11	620
	≥ 144	5	19	1094
BUN	<11	0	0	0
	[11, 15)	0	0	3
	[15, 19)	0	0	7
	[19, 27)	0	0	14
	≥ 27	0	0	24
Anion gap	<5	3	11	648
	[5, 7)	3	10	575
	[7, 10)	2	8	453
	[10, 12)	2	6	330
	[12, 17)	1	3	159
	≥ 17	0	0	0
Race	African American	1	4	248
	Other	1	2	124
	Caucasian	0	0	0

Variable	Level	Risk score for HFR		
		B^*	$5\beta_{age}$	B_{min}
BUN	<11	0	0	0
	[11, 15)	1	4	727
	[15, 19)	1	6	1308
	[19, 27)	2	11	2180

(continued)

Table 3. (continued)

Variable	Level	Risk score for HFR		
		B^*	$5\beta_{age}$	B_{min}
Hemoglobin	≥ 27	3	19	3778
	<9	1	5	945
	[9, 9.7)	1	3	702
	[9.7, 10.3)	0	3	516
Hematocrit	[10.3, 11.1)	0	2	315
	≥ 11.1	0	0	0
	<26.9	0	0	0
	[26.9, 28.8)	0	2	377
	[28.8, 30.7)	1	3	682
	[30.7, 33.1)	1	5	1028
Length of stay	≥ 33.1	1	8	1573
	<5	0	0	0
	[5, 7)	1	4	811
	[7, 14)	2	13	2637
preInp1Y ^a	≥ 14	4	24	4869
	0	0	0	0
preER1Y ^b	1	1	6	1300
	≥ 2	2	13	2600
	0	0	0	0
Charlson Comorbidity Index	1	0	2	475
	≥ 2	1	5	949
	<4	0	0	0
	[4, 6)	1	8	1678
Age	≥ 6	2	10	2098
	<65	0	0	0
	[65, 75)	0	2	443
	[75, 80)	1	4	745
	[80, 85)	1	5	946
Platelet count	[85, 90)	1	6	1147
	≥ 90	1	6	1248
	<143	0	0	0
	[143, 177)	0	0	39
	[177, 213)	0	0	74
Potassium	[213, 268)	1	1	120
	≥ 268	1	1	200
	<3.6	0	1	260
	[3.6, 3.9)	0	1	184
Sodium	[3.9, 4.1)	0	1	130
	[4.1, 4.4)	0	0	76
	≥ 4.4	0	0	0
	<134	0	1	112
Creatinine	[134, 136)	0	0	75
	[136, 138)	0	0	50
	[138, 140)	0	0	25
	≥ 140	0	0	0
	<0.6	0	0	92
Creatinine	[0.6, 0.8)	0	0	75
	[0.8, 0.9)	0	0	61
	[0.9, 1.2)	0	0	39
	≥ 1.2	0	0	0

^a Number of inpatient visits within 1 year before.

^b Number of emergency department visits within 1 year before.

multiple accuracy measures while streamlining the risk score scale presents an intriguing avenue for future work.

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Author contributions

Yajun Lu (Conceptualization, Data curation, Methodology, Coding, Formal analysis, Writing—original draft, Writing—review & editing), Thanh Duong (Data curation, Coding, Formal analysis), Zhuqi Miao (Conceptualization, Data curation, Methodology, Coding, Formal analysis, Writing—original draft, Writing—review & editing), Thanh Thieu (Conceptualization, Methodology, Formal analysis), Jivan Lamichhane (Writing—review & editing, Validation), Abdulaziz Ahmed (Writing—review & editing, Validation), and Dursun Delen (Conceptualization, Writing—review & editing).

Supplementary material

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Conflicts of interest

None declared.

Data availability

The data used in this study were obtained from Oracle Cerner. Interested researchers may request the associated data directly from Oracle Cerner.

Code availability

The code supporting the findings of this study is shared publicly on GitHub at <https://github.com/yajun668/RiskScoring>.

Ethics information

The Institutional Review Boards (IRB) at Oklahoma State University exempted the study from review because the data have been completely de-identified according to HIPAA regulations. The entire process of data collection and analysis took place on devices associated with Oklahoma State University.

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